Spotlights on the theory of elicitability Habilitation Talk

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Overview of Habilitation thesis

Publications included in Habilitation thesis

- T. Fissler, S. M. Pesenti (2023). Sensitivity measures based on scoring functions. European Journal of Operational Research 307 (3), 1408–1423
- T. Fissler, H. Holzmann (2022). Measurability of functionals and of ideal point forecasts. *Electronic Journal of Statistics* 16 (2), 5019–5034
- C. Heinrich-Mertsching, T. Fissler (2022). Is the mode elicitable relative to unimodal distributions? *Biometrika* 109 (4), 1157–1164
- T. Fissler, J. F. Ziegel (2021). On the elicitability of range value at risk. Statistics & Risk Modeling 25 (1-2), 25-46
- 5. T. Fissler, J. Hlavinová, B. Rudloff (2021). Elicitability and Identifiability of Systemic Risk Measures. *Finance and Stochastics* **25** (1), 133–165
- T. Fissler, R. Frongillo, J. Hlavinová, B. Rudloff (2021). Forecast evaluation of quantiles, prediction intervals, and other set-valued functionals.

Electronic Journal of Statistics 15 (1), 1034–1084

Publications not included in the thesis

- T. Fissler, Y. Hoga (2023). Backtesting Systemic Risk Forecasts using Multi-Objective Elicitability. Journal of Business & Economic Statistics.
- 8. T. Dimitriadis, T. Fissler, J. Ziegel (2023). Characterizing M-estimators. *Biometrika (forthcoming).*
- 9. T. Dimitriadis, T. Fissler, J. Ziegel (2023). Osband's Principle for Identification Functions. *Statistical Papers*
- T. Fissler, M. Merz, M. V. Wüthrich (2023). Deep Quantile and Deep Composite Model Regression. *Insurance: Mathematics and Economics* 109, 94–112
- T. Fissler, J. F. Ziegel (2019). Order-Sensitivity and Equivariance of Scoring Functions. *Electronic Journal of Statistics* 13 (1), 1166–1211
- T. Fissler, M. Podolskij (2017). Testing the maximal rank of the volatility process for continuous diffusions observed with noise.

Bernoulli 23 (4B), 3021-3066

Publications not included in the thesis

- T. Fissler, J. F. Ziegel (2016). Higher order elicitability and Osband's principle. Annals of Statistics 44 (4), 1680–1707.
- T. Fissler, J. F. Ziegel, T. Gneiting (2016). Expected Shortfall is jointly elicitable with Value at Risk – Implications for backtesting. *Risk Magazine*, January 2016, 58–61.
- T. Fissler, C. Thäle (2016).
 A four moments theorem for Gamma limits on a Poisson chaos.
 ALEA, Lat. Am. J. Probab. Math. Stat. 13 (1), 163–192.

Preprints

- T. Fissler, C. Lorentzen, M. Mayer (2022). Model Comparison and Calibration Assessment: User Guide for Consistent Scoring Functions in Machine Learning and Actuarial Practice. https://doi.org/10.48550/arXiv.2202.12780
- T. Dimitriadis, T. Fissler, J. F. Ziegel (2020). The Efficiency Gap. https://doi.org/10.48550/arXiv.2010.14146

Roadmap

- 1. Overview of Habilitation thesis \checkmark
- 2. Setup
- 3. Loss functions in statistical learning & forecast comparison
- 4. The elicitation problem
- 5. The wondrous tale about the mode
- 6. Range Value at Risk a linear combination of Bayes risks
- 7. The dichotomy of the set-valued elicitation world
- 8. Summary and outlook

Setup

Setup

- Y: The quantity of interest or response:
 - Typically real-valued, but could also be multivariate, categorical etc.
 - Examples: Claim sizes, number of claims, temperature, precipitation, wind speed, demand for a product, GDP growth, inflation, loss of a company
- X: Explanatory variables, regressors, features:
 - From a possibly high dimensional feature space \mathcal{X} .
 - · Can contain metrical variables, categorical etc.
 - Can be exogenous variables (cross-sectional), but also past observations of Y (time series setup)

Learning We want to exploit the information in \boldsymbol{X} to describe Y as accurately as possible.

 \sim How to fit a model?

Prediction We want to exploit the information in X to predict unseen Y as accurately as possible.

 \sim How to assess the accuracy?

Define your goal!

- Usually, **X** does not fully describe Y: There is no deterministic function g such that $Y = g(\mathbf{X})$.
- The remaining uncertainty of Y given **X** can be described in terms of the conditional distribution

 $F_{Y|X}$

Define your goal!

- Probabilistic predictions: Try to learn the full conditional distribution and come up with probabilistic forecasts $\hat{F}_{Y|X}$.
 - Very informative approach.
 - Often hard to implement.
 - Can be difficult to communicate.
- Point predictions: Summarise the conditional distribution with a functional of the conditional distribution

$$T(Y \mid \boldsymbol{X}) := T(F_{Y \mid \boldsymbol{X}})$$

Note: The existence and $\sigma(\mathbf{X})$ -measurability of $T(Y \mid \mathbf{X})$ has been established in F. & Holzmann (EJS, 2022).

• Examples:

- mean, median, mode
- quantiles, expectiles
- Risk measures: Value at Risk, Expected Shortfall
- Come up with point forecasts $\widehat{T}(Y \mid X)$.
 - Loss of information
 - Easier to implement
 - Easier to communicate.

Loss functions in statistical learning & forecast comparison

Action domain and model choice

- Let \mathcal{F} be a convex class of distributions such that $F_{Y|X} \in \mathcal{F}$.
- Call the space where the chosen target functional maps to action domain A. T: F → A.
- Examples:
 - + $\mathcal{A}=\mathbb{R}$ for the mean / quantile or $[0,\infty)$ for the mean / quantile of a positive Y
 - + $\mathcal{A}=\mathbb{R}^k$ for mean of a multivariate observation, different quantiles of a real-valued observation
 - $\blacktriangleright \ \mathcal{A}$ finite for the mode of a categorical observation
 - + $\mathcal{A} \subseteq \mathcal{P}(\mathbb{R}^k)$ for prediction sets or systemic risk measures.
 - A = F, a class of probability distributions or densities for probabilistic forecasts (then T is the identity functional).
- Consider a model class \mathcal{M} of models $m: \mathcal{X} \to \mathcal{A}$.
- Examples:
 - (Generalised) Linear Models
 - Neural nets
 - Isotonic regression functions
- Convexity of \mathcal{F} ensures that $F_Y, F_{Y|m(\mathbf{X})} \in \mathcal{F}$ for all $m \in \mathcal{M}$.

Consistent loss functions and elicitability

Definition 1 (Consistency)

A loss function is a map

$$L\colon \mathcal{A}\times\mathbb{R}\to\mathbb{R}.$$

Sometimes, additional assumptions are imposed such as continuity (in the first argument), positivity etc. It is \mathcal{F} -consistent for a functional T if

$$\mathrm{E}_{Y \sim F}\left[L(\mathit{T}(\mathit{F}), \mathit{Y})\right] \leqslant \mathrm{E}_{Y \sim F}\left[L(\mathit{a}, \mathit{Y})\right] \qquad \text{for all } \mathit{a} \in \mathcal{A}, \ \mathit{F} \in \mathcal{F}.$$

L is strictly \mathcal{F} -consistent if equality arises only if a = T(F).

Definition 2 (Elicitability)

A functional T is elicitable on F if there is a strictly F-consistent loss function for it.

Alternative name for loss functions: Scoring functions

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First examples of elicitable functionals

The mean is elicitable on the class of square integrable distributions. A strictly consistent loss function is given via the squared loss

$$L(a, y) = (a - y)^2.$$

The α -quantile is elicitable on the class of integrable distributions which are strictly increasing. A strictly consistent loss function is given via the pinball loss / asymmetric piecewise linear loss

$$L(\mathbf{a}, \mathbf{y}) = (\mathbb{1}\{\mathbf{y} \leq \mathbf{a}\} - \alpha)(\mathbf{a} - \mathbf{y}).$$

Learning via loss minimisation (M-estimation)

• Consider the statistical risk

$$R(m) = E \left[L(m(\boldsymbol{X}), Y) \right]$$
$$= E \left[E \left[L(m(\boldsymbol{X}), Y) \right] | \boldsymbol{X} \right]$$

• Bayes rule is given by

$$m^* \in \underset{m \in \mathcal{M}}{\operatorname{arg\,min}} R(m).$$

• If the true regression function $\mathbf{x} \mapsto T(Y \mid \mathbf{X} = \mathbf{x})$ is in \mathcal{M} and if L is \mathcal{F} -consistent for T, we get

$$\mathbb{E}\left[L\big(\mathcal{T}(\mathcal{Y} \mid \boldsymbol{X}), \mathcal{Y})\big) \mid \boldsymbol{X}\right] \leq \mathbb{E}\left[L\big(m(\boldsymbol{X}), \mathcal{Y})\big) \mid \boldsymbol{X}\right].$$

- Therefore, $T(Y | X = \cdot)$ is a Bayes rule.
- Due to Dimitriadis, F., Ziegel (Biometrika, 2023), the (strict) consistency of *L* is also necessary for $T(Y | X = \cdot)$ to be the only Bayes act (under certain richness conditions).

Learning via loss minimisation (M-estimation)

• Let $D_{\text{train}} = \{(\mathbf{x}_i, y_i), i = 1, ..., n\}$ be a training sample. Define the empirical risk

$$\overline{R}(m; D_{\text{train}}) = \frac{1}{n} \sum_{(\mathbf{x}_i, y_i) \in D_{\text{train}}} L(m(\mathbf{x}_i), y_i)$$
$$\approx \mathbb{E} \left[L(m(\mathbf{X}), Y) \right]$$
$$= R(m).$$

• M-estimator \hat{m} is an empirical risk minimiser

$$\widehat{m} \in \operatorname*{arg\,min}_{m \in \mathcal{M}} \overline{R}(m; D_{\mathsf{train}})$$

Pitfall of overfitting

• Estimator \hat{m} depends on training sample D_{train} :

- Prone to estimation error
- Different training samples lead to different estimates.
- Danger that \hat{m} learns the noisy pattern of the sample at hand and not the structure of the distribution.
- In-sample performance $\overline{R}(\hat{m}; D_{\text{train}})$ can be a bad estimate for the actual risk $R(\hat{m})$.
- \rightarrow pitfall of overfitting.
- This problem gets bigger
 - the more complex a model is;
 - * the smaller (less representative) the trainings sample is.

Pitfall of overfitting



Mitigating overfitting

There are two main strategies:

1. Fitting: Introduce a penalty term Ω , accounting for the model complexity:

$$\widehat{m} = \operatorname*{arg\,min}_{m \in \mathcal{M}} \overline{R}(m; \mathcal{D}_{\mathsf{train}}) + \lambda \Omega(m).$$

Examples for Ω :

- \blacktriangleright Number of parameters \rightsquigarrow AIC and BIC
- \blacktriangleright Norms of the parameter \leadsto ridge and lasso regression
- Number of optimisation steps when fitting a neural net
- 2. Validation: Monitor the out-of-sample risk on an (ideally) independent and identically distributed validation set $D_{\text{valid}} = \{(\mathbf{x}_i, y_i), i = 1, \dots, l\}$ via

$$\overline{R}(\widehat{m}; D_{\mathsf{valid}})$$

- Better approximation of the statistical risk.
- Can be made more efficient with cross-validation.

Model agnostic forecast comparison

- Suppose the target functional T is fixed (could also be probabilistic).
- We have different methods of producing predictions, but we are agnostic about how they have been produced.
- \sim We adhere to the weak prequential principle (Dawid & Vovk, 1999).
 - Example for two different forecasters: We have the prediction-observation sequence

$$(A_i^{(1)}, A_i^{(2)}, Y_i)$$
 $i = 1, \dots, n$

• Ranking in terms of the empirical loss difference

$$\frac{1}{n}\sum_{i=1}^{n} L(A_{i}^{(1)}, Y_{i}) - L(A_{i}^{(2)}, Y_{i})$$

- Forecast method 1 is deemed better than 2 if this is negative.
- Tests for equal predictive accuracy E[L(A⁽¹⁾, Y)] = E[L(A⁽²⁾, Y)] and forecast dominance E[L(A⁽¹⁾, Y)] ≥ E[L(A⁽²⁾, Y)] can be assessed via Diebold–Mariano tests (amounting to *t*-tests).
- To honour truthful forecasting, L should be (strictly) consistent for T!

The Elicitation Problem

The Elicitation Problem

- Fix some functional $T: \mathcal{F} \to \mathcal{A}$.
- (a) Is T elicitable?
- (b) What is the class of (strictly) consistent loss functions for T?
- (c) What is a particularly good choice of a loss function?
- (d) What to do if T is not elicitable?

| <i>T</i> | L(x, y) |
|-------------------------------------|---|
| mean | $(x-y)^2$ |
| median | x-y |
| au-expectile | $ \mathbb{1}\{y \le x\} - \tau (x - y)^2$ |
| lpha-quantile | $ \mathbb{1}\{y \leq x\} - \alpha x - y $ |
| variance | × |
| Expected Shortfall | × |
| (mean, variance) | \checkmark |
| (Value at Risk, Expected Shortfall) | \checkmark |
| identity (probabilistic forecast) | $L(F, y) = -\log(f(y))$ |

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(a) One-dimensional functionals

Theorem 3 (Convex level sets, Osband, 1985) Let $T: \mathcal{F} \to \mathcal{A}$ be an elicitable functional. Let $F_0, F_1 \in \mathcal{F}$ and $F_{\lambda} = (1 - \lambda)F_0 + \lambda F_1 \in \mathcal{F}$ for some $\lambda \in (0, 1)$. Then

$$T(F_0) = T(F_1) \implies T(F_\lambda) = T(F_0)$$

Proof: Let $t = T(F_0) = T(F_1)$ and $x \neq t$. Then, due to the linearity of the expectation in the measure,

$$\begin{split} \mathbf{E}_{Y \sim F_{\lambda}}[L(t, Y)] &= (1 - \lambda) \, \mathbf{E}_{Y \sim F_{0}}[L(t, Y)] + \lambda \, \mathbf{E}_{Y \sim F_{1}}[L(t, Y)] \\ &\leq (1 - \lambda) \, \mathbf{E}_{Y \sim F_{0}}[L(x, Y)] + \lambda \, \mathbf{E}_{Y \sim F_{1}}[L(x, Y)] \\ &= \mathbf{E}_{Y \sim F_{\lambda}}[L(x, Y)]. \end{split}$$

(a) One-dimensional functionals

Theorem 4 (Convex level sets, Osband, 1985) Let $T: \mathcal{F} \to \mathcal{A}$ be an elicitable functional. Let $F_0, F_1 \in \mathcal{F}$ and $F_{\lambda} = (1 - \lambda)F_0 + \lambda F_1 \in \mathcal{F}$ for some $\lambda \in (0, 1)$. Then

$$T(F_0) = T(F_1) \implies T(F_\lambda) = T(F_0)$$

Remarks:

• This shows that the variance or ES are generally not elicitable.

$$\operatorname{Var}(\delta_{\mathsf{x}}) = \operatorname{Var}(\delta_{\mathsf{y}}) = 0, \quad \operatorname{Var}(\lambda\delta_{\mathsf{x}} + (1-\lambda)\delta_{\mathsf{y}}) = \lambda(1-\lambda)(\mathsf{x}-\mathsf{y})^2.$$

- Steinwart et al. (2014) showed that for A = ℝ and under some continuity assumptions on T, CxLS are also sufficient for elicitability.
- This argument is independent of the dimension of *T*.
- For k > 1, CxLS are generally not sufficient, e.g., $(VaR_{\alpha}, CoVaR_{\alpha|\beta})$.

The wondrous tale about the mode

The wondrous tale about the mode

- The mode is the argmax of the counting / Lebesgue density.
- The mode functional has CxLS.
- On classes of discrete distributions only (say on ℕ), it is elicitable with the zero-one loss:

$$L(x, y) = \mathbb{1}\{x \neq y\}$$

• What about absolutely continuous distributions? Clearly, the zero-one loss is constant almost surely. But are there other candidates?

Theorem 5 (Heinrich-Mertsching and F. (Biometrika, 2022))

The mode is not elicitable on \mathcal{F}_0 , the class of continuous and strongly unimodal densities on \mathbb{R} .

(And hence it fails to be elicitable on any superclass \mathcal{F} of $\mathcal{F}_{0.}$)

This result substantially strengthens the result of Heinrich (2014) which establishes the non-elicitability on the class containing *all* absolutely continuous distributions on \mathbb{R} .

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The wondrous tale about the mode (continued)

Proof:

• Key observation: The mode fails to be continuous:

For any a < b there are sequences of densities $f_n, g_n \in \mathcal{F}_0$

- ▶ $\operatorname{mode}(f_n) =: x_1 \neq x_2 := \operatorname{mode}(g_n) \text{ for all } n \in \mathbb{N}, x_1, x_2 \in [a, b];$
- f_n and g_n converge pointwise to the uniform distribution on [a, b] (which is not contained in \mathcal{F}_0).

• If a strictly \mathcal{F}_0 -consistent loss function L existed, this would imply that

$$\int_{a}^{b} L(x_1, y) \mathrm{d}y = \int_{a}^{b} L(x_2, y) \mathrm{d}y \tag{1}$$

 Since (1) holds for all a < b, the Radon–Nikodym theorem implies that

$$L(x_1, y) = L(x_2, y) \text{ for almost all } y.$$
(2)

• (2) shows that L cannot be strictly \mathcal{F}_0 -consistent.

Range Value at Risk – a linear combination of Bayes risks

What to do if T is not elicitable?

A lot of (relevant) functionals are not elicitable: Variance, Expected Shortfall (ES), but also Range Value at Risk (RVaR).

$$ES_{\alpha}(Y) = \frac{1}{\alpha} \int_{\alpha}^{1} VaR_{\gamma}(Y)d\gamma$$

$$\stackrel{(\star)}{=} E\left[Y \mid VaR_{\alpha}(Y) \leq Y\right]$$

$$RVaR_{\alpha,\beta}(Y) = \frac{1}{\beta - \alpha} \int_{\alpha}^{\beta} VaR_{\gamma}(Y)d\gamma$$

$$\stackrel{(\star)}{=} E\left[Y \mid VaR_{\alpha}(Y) \leq Y \leq VaR_{\beta}(Y)\right]$$

- $RVaR_{\alpha,\beta}$ is an interpolation of ES_{α} and VaR_{α} .
- It is robust, but not coherent.
- Moreover, $RVaR_{\alpha,1-\alpha}$ is a trimmed mean.

What to do if T is not elicitable?

Variance and ES can be written as the Bayes risk of a consistent loss function.

$$Var(Y) = \min_{x \in \mathbb{R}} E\left[(x - Y)^2\right]$$
$$ES_{\alpha}(Y) = \min_{x \in \mathbb{R}} E\left[\frac{1}{\alpha}S_{\alpha}(x, Y)\right], \quad S_{\alpha}(x, y) = (\mathbb{1}\{y \le x\} - \alpha)x - \mathbb{1}\{y \le x\}y$$

Theorem 6

Let T be elicitable with strictly consistent loss S. Then (T, T^*) is jointly elicitable where

$$T^{\star}(F) = \min_{x \in \mathcal{A}} \mathbb{E}_{Y \sim F} \left[S(x, Y) \right].$$

A strictly consistent loss for $(\mathit{T}, \mathit{T}^\star)$ is given by

$$L(x_1, x_2; y) = \phi'(x_2) (x_2 - S(x_1, y)) - \phi(x_2) + L_T(x_1, y),$$

where ϕ is strictly convex, $\phi' < 0$, and L_T is a consistent loss for T.

$$L(x_1, x_2; y) = \phi'(x_2) \big(x_2 - S(x_1, y) \big) - \phi(x_2) + L_T(x_1, y),$$

Idea:

• For fixed x₂, the map

$$(x_1, y) \mapsto L(x_1, x_2; y) = -\phi'(x_2)S(x_1, y) + L_T(x_1, y) + \kappa(x_2)$$

is strictly consistent for *T*, since $\phi' < 0$.

• For fixed x₁, the map

$$(x_2, y) \mapsto L(x_1, x_2; y) = \phi'(x_2) (x_2 - S(x_1, y)) - \phi(x_2) + \dots$$

is strictly consistent for $F \mapsto \operatorname{E}_{Y \sim F} S(x_1, Y)$, since ϕ is convex.

Corollary 7 (F and Ziegel (AoS, 2016)) The pairs (mean, variance) and $(VaR_{\alpha}, ES_{\alpha})$ are elicitable!

Elicitability of RVaR

RVaR is the scaled difference of Bayes risks!

$$\mathrm{ES}_{\alpha}(\mathbf{Y}) = \frac{1}{\beta - \alpha} \left(\min_{\mathbf{x} \in \mathbb{R}} \mathrm{E} \left[\mathbf{S}_{\alpha}(\mathbf{x}, \mathbf{Y}) \right] - \min_{\mathbf{x} \in \mathbb{R}} \mathrm{E} \left[\mathbf{S}_{\beta}(\mathbf{x}, \mathbf{Y}) \right] \right)$$

Theorem 8 (F and Ziegel (Stat. Risk Model., 2021))

$$L(x_1, x_2, x_3; y) = (\mathbb{1}\{y \le x_1\} - \alpha)(g_1(x_1) - g_1(y))$$

$$+ (\mathbb{1}\{y \le x_2\} - \beta)(g_2(x_2) - g_2(y))$$

$$+ \phi'(x_3) \left(x_3 - \frac{1}{\beta - \alpha} \left(S_\alpha(x_1, y) - S_\beta(x_2, y)\right)\right) - \phi(x_3)$$
(3)

is strictly consistent for $(\mathrm{VaR}_{\alpha},\mathrm{VaR}_{\beta},\mathrm{RVaR}_{\alpha,\beta})$ if

- ϕ is strictly convex;
- for all x_3 : $x_1 \mapsto g_1(x_1) x_1 \phi'(x_3)/(\beta \alpha)$ is strictly increasing;
- for all $x_3: x_2 \mapsto g_2(x_2) + x_2 \phi'(x_3)/(\beta \alpha)$ is strictly increasing.

Any strictly consistent loss is essentially of the form (3).

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The dichotomy of the set-valued elicitation world

Set-valued functionals

The functional T maps to a subset of $\mathcal{P}(\mathbb{R}^k)$.

- Mode: $mode(F) = argmax_x f(x)$
- Quantiles: $q_{\alpha}(F) := \{x \in \mathbb{R} \mid \lim_{t \uparrow x} F(t) \leq \alpha \leq F(x)\}$
- Prediction intervals: Any [a, b] s.t. $F([a, b]) := F(b) F(a-) \ge \alpha$
- Systemic risk measures: R(F_Y) = {k ∈ ℝ^d | ρ(Λ(Y + k)) ≤ 0}, see Feinstein, Rudloff and Weber (2017)
- Functionals of random sets: Climatology, reliability engineering, medicine, econometrics; see Molchanov (2017); Molchanov and Molinari (2018).

Selective vs. exhaustive forecasts

• Example of the α -quantile $q_{\alpha} \colon \mathbb{R} \to \mathcal{P}(\mathbb{R})$

$$q_{\alpha}(F) = \{ x \in \mathbb{R} \mid \lim_{t \uparrow x} F(t) \leq \alpha \leq F(x) \} \subset \mathbb{R}.$$

• Choice of the action domain \mathcal{A} :

 $\mathcal{A}_{sel} \subseteq \mathbb{R}$: The forecasts are points in \mathbb{R} . There are multiple best actions, namely each selection $x \in q_{\alpha}(F)$. \sim Selective forecasts

 $\mathcal{A}_{exh} \subseteq \mathcal{P}(\mathbb{R})$: The forecasts are subsets of \mathbb{R} . There is a unique best action namely to report the entire set $B = q_{\alpha}(F)$. \sim Exhaustive forecasts

Two modes of elicitability

Definition 9 (Elicitability)

(a) A functional $T: \mathcal{F} \to \mathcal{P}(\mathcal{A}_{sel})$ is selectively elicitable if there is a selective loss function $L_{sel}: \mathcal{A}_{sel} \times O \to \mathbb{R}$ such that

$$\mathrm{E}_{F}[L_{\mathrm{sel}}(t, Y)] \leq \mathrm{E}_{F}[L_{\mathrm{sel}}(x, Y)]$$

for all $F \in \mathcal{F}$, for all $t \in T(F)$, for all $x \in \mathcal{A}_{sel}$ and where equality implies that $x \in T(F)$.

(b) A functional $T: \mathcal{F} \to \mathcal{A}_{exh}$ is exhaustively elicitable if there is an exhaustive loss function $\mathcal{L}_{exh}: \mathcal{A}_{exh} \times O \to \mathbb{R}$ such that

$$\mathbb{E}_{F}[L_{\text{exh}}(T(F), Y)] \leq \mathbb{E}_{F}[L_{\text{exh}}(B, Y)]$$

for all $F \in \mathcal{F}$, for all $B \in \mathcal{A}_{exh}$ and where equality implies that B = T(F).

Mutual exclusivity results

Theorem 10 (F., Frongillo, Hlavinová, Rudloff (EJS, 2021))

- If there exist $F, G \in \mathcal{F}$ such that $\emptyset \neq T(F_0) \subsetneq T(F_1)$ and $(1-\lambda)F_0 + \lambda F_1 \in \mathcal{F}$ for all $\lambda \in (0,1)$ then:
 - (i) If T is selectively elicitable, it is not exhaustively elicitable.
 - (ii) If T is exhaustively elicitable, it is not selectively elicitable.

Main idea for proofs:

Exploit "linearity" of the expected loss $\overline{L}(x, F)$ in its second argument and use a refinement of the fact that convex level sets are necessary for elicitability.

Examples: Mode and Quantiles

• If all $F \in \mathcal{F}_{count}$ have countable support, the mode is selectively elicitable on \mathcal{F}_{count} with the loss

$$L_{mode}(x, y) = \mathbb{1}\{x \neq y\}.$$

 \rightsquigarrow The mode is generally not exhaustively elicitable.

• The α -quantile is selectively elicitable with a strictly consistent loss

$$L_{\alpha}(x, y) = |\mathbb{1}\{y \leq x\} - \alpha ||x - y|.$$

 \rightsquigarrow The $\alpha\text{-quantile}$ is generally not exhaustively elicitable.

• One can also show that the lower quantile (or any other selection of it) is in general not elicitable!

Examples: Prediction intervals and systemic risk measures

- The class of α -prediction intervals is exhaustively elicitable.
- The class of shortest α -prediction interval is neither selectively nor exhaustively elicitable.
- Systemic risk measures of the form $R(F_Y) = \{k \in \mathbb{R}^d | \rho(\Lambda(Y+k)) \leq 0\}$ are exhaustively elicitable, if ρ is elicitable. (F, Hlavinová, Rudloff (Fin. Stoch., 2021)).

Closed Random Sets

- Let $(\Omega, \mathfrak{F}, \mathbf{P})$ be a non-atomic probability space.
- A closed random set \mathbf{Y} is a map from Ω into the collection \mathfrak{U} of closed sets in \mathbb{R}^d (or some general separable Banach space).
- It is measurable if for all compact sets $K \subseteq \mathbb{R}^d$

$$\{\omega \mid \mathbf{Y}(\omega) \cap \mathbf{K} \neq \emptyset\} \in \mathfrak{F}.$$

See Molchanov (2017) for details.

- Examples:
 - region of a blackout in a country
 - region affected by a flood, avalanche, disease
 - tumorous tissue in the human body
- There are interesting set-valued functionals of random sets:
 - Vorob'ev quantiles
 - Vorob'ev expectation
 - Selection expectation (\approx Minkowski average)

Vorob'ev Quantiles of Closed Random Sets

Definition 11

The upper excursion set of the coverage function $u \mapsto \mathbf{P}(u \in \mathbf{Y})$ at level $\alpha \in [0, 1]$,

$$Q_{\alpha}(\mathbf{Y}) := \{ u \in \mathbb{R}^d \, | \, \mathbf{P}(u \in \mathbf{Y}) \ge \alpha \},\$$

is called the Vorob'ev α -quantile of **Y**.

 $Q_{\alpha}(\mathbf{Y})$ is always a closed set.

Theorem 12 (F., Frongillo, Hlavinová, Rudloff (EJS, 2021))

(i)

$$L: \mathfrak{U} \times \mathfrak{U} \to [0, \infty], \quad L(X, Y) = \alpha \mu(X \setminus Y) + (1 - \alpha) \mu(Y \setminus X),$$

is a non-negative ${\cal F}$ -consistent loss function for ${\it Q}_{lpha}/$

(ii) If Q_α(F) = cl(Q_α[>](F)) and Q_α(F) = cl(int(Q_α(F)) for all F ∈ F, then Q_α is exhaustively elicitable on F. Moreover, for any σ-finite positive measure μ on ℝ^d such that E_F[μ(Y)] < ∞ and π(Q_α(F)) < ∞ for all F ∈ F, the restriction of L to the family U': = {U ∈ U | U = cl(int(U))} is a strictly F-consistent exhaustive loss function for Q_α.

Tobias Fissler (ETH Zurich)

Interpretation of loss

$$L(X, Y) = \alpha \mu(X \setminus Y) + (1 - \alpha) \mu(Y \setminus X)$$

Decomposition into

- false positive $X \setminus Y$
- false negative $Y \setminus X$

Applications:

- Evaluation of warnings (in spacetime) where asymmetric costs for false positives and false negatives are present.
- Pattern recognition in learning and diagnostics.
- Mathematical statistics: A confidence set is actually a random set.

Summary

- Loss functions play a crucial role learning and in forecast assessment and comparison.
- They should be chosen in line with the target functional of interest. \sim They should be consistent.
- Strict consistency ensures that the oracle regression function is eventually learned. It ensures incentive compatible forecast comparison.
- We have revisited the elicitation problem.
 - CxLS are necessary for elicitability. The mode shows that they are not generally sufficient.
 - Linear combinations of Bayes risks are elicitable.
 - A refinement of the CxLS property establishes that set-valued functionals can either be selectively elicitable, exhaustively elicitable or not elicitable at all.

Omitted achievements, open questions, outlook

• Omitted achievements:

- Discussion of calibration assessment with identification functions.
- Loss functions as measures of information (generalising the coefficient of determination) (F & Pesenti, EJOR, 2023)
- Multivariate loss functions (F & Hoga, JBES, 2023)
- Loss functions in modern statistical learning (= machine learning) (F, Merz, & Wüthrich, IME, 2023)

• Open questions & outlook:

- Replace strict consistency (ranking of expectations) by requirement that average scores rank with high probability.
- Better understanding of generative AI such as ChatGPT. ("Small" input vector / prompt associated with very complex response)

Further Reading

• Scoring rules for probabilistic forecasts:

T. Gneiting and A. E. Raftery. Strictly proper scoring rules, prediction, and estimation.

Journal of the American Statistical Association, 102:359-378, 2007

• Good introduction to elicitability:

T. Gneiting. Making and evaluating point forecasts.

Journal of the American Statistical Association, 106(494):746-762, 2011

• Traditional and Comparative backtests:

N. Nolde and J. F. Ziegel. Elicitability and backtesting: Perspectives for banking regulation.

The Annals of Applied Statistics, 11(4):1833-1874, 2017

Additional References

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- Z. Feinstein, B. Rudloff, and S. Weber. Measures of systemic risk. *SIAM Journal on Financial Mathematics*, 8:672–708, 2017
- C. Heinrich. The mode functional is not elicitable. *Biometrika*, 101(1):245–251, 2014
- I. Molchanov. *Theory of Random Sets*.
 Probability Theory and Stochastic Modelling. Springer-Verlag London, London, 2 edition, 2017
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- I. Steinwart, C. Pasin, R. Williamson, and S. Zhang. Elicitation and Identification of Properties. JMLR Workshop Conf. Proc., 35:1–45, 2014

Thank you for your attention! Looking forward to our discussion!