Risk simulation with optimally stratified importance sampling

W. Hörmann and İ. Başoğlu
Bogazici University Istanbul
Industrial Engineering Department

June 12, 2014

Some Risk Measures

- Tail loss probability for given threshold level, τ:
  $\Pr(Loss \geq \tau) = E\left[\mathbb{1}\left\{Loss \geq \tau\right\}\right]$
- $VaR(\alpha)$: Value-at-risk, $1 - \alpha$ quantile of the loss distribution.
- Conditional excess: $E\left[Loss | Loss \geq \tau\right]$
- Conditional value-at-risk: $CVaR(\alpha)$: $E\left[Loss | Loss \geq VaR(\alpha)\right]$

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t-copula for modeling log-returns

- For modeling dependence among log-returns the importance of copulae is stressed (Frey and McNeil, 2001).
- t-copula models have a good fit to the joint distribution of log returns. (for example Mashal et al. 2003; Kole et al. 2007).

$$Return\,(\mathbf{T}) = \sum_{d=1}^{D} w_d e^{c_d G_d^{-1}(F_\nu(T_d))}$$

$$c_d = \sqrt{\frac{\sigma_d^2}{252}\frac{1}{\mathrm{var}_d}}$$

$$\boldsymbol{T} = (T_1, \ldots, T_D)' = \frac{\boldsymbol{LZ}}{\sqrt{Y/\nu}}$$

Our Objectives

For a linear asset portfolio of moderate size (2 to 10 assets):

- Efficient estimation of a single tail loss probability, $\Pr(Loss \geq \tau)$
- Efficient estimation of a single conditional excess $E\left[Loss|Loss \geq \tau\right]$.
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- Estimation of an expectation: $x = E_f[q(\boldsymbol{X})]$
  - $\boldsymbol{X} \in \mathbb{R}^D$ and $\boldsymbol{X}$ has density $f(.)$,
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  - Error bound: $\hat{x}_{NV} \pm \Phi^{-1}(\alpha/2)\, \sigma \big/ \sqrt{N}$.
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Importance Sampling (IS)

IS is a frequently used method for rare event situations

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  That density is unknown for relevant applications.
- In practice an IS density is typically taken from a parametric family (often the same as $f(.)$). The parameters are selected such that the IS density imitates $\left|q\left(\boldsymbol{x}\right) f\left(\boldsymbol{x}\right)\right|$.
- The cross entropy method is a general approach to select the parameters of the IS density.
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  That density is unknown for relevant applications.

- In practice an IS density is typically taken from a parametric family (often the same as $f(.)$). The parameters are selected such that the IS density imitates $\left|q\left(\boldsymbol{x}\right) f\left(\boldsymbol{x}\right)\right|$.
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Sak, WH and Leydold (2010) use IS for the iid normal input Z and the chi-square random variate Y.

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- Let ξ_i, $i = 1, \ldots, I$ be a partition of $\mathbb{R}^D$ into I strata,
- Assume $p_i = \Pr\{\boldsymbol{X} \in \xi_i\}$ are known for $i = 1, \ldots, I$.
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- Define allocation fractions $\pi_i = N_i / N$

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- The Adaptive Optimal Allocation (AOA) algorithm terminates in K iterations.
- The total sample size N is divided between iterations with a non-decreasing order (e.g., $K = 3$, $0.1N, 0.4N, 0.5N$).
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- OAS can be interpreted as IS using the product of the original density and a step function as weight function.
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- High dimensional stratification not possible in practice. Thus (like for IS) one (or two) main directions are used for most applications.
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- Disadvantage of OAS: Many strata (or an adaptive strata structure) are necessary for rare event simulations.
- Is it possible and sensble to combine IS and OAS to increase the variance reduction?

IS vs Optimal Allocation Stratification (OAS)

- OAS can be interpreted as IS using the product of the original density and a step function as weight function.
- Advantage of OAS is the simple formula for the optimal allocation fractions.
- High dimensional stratification not possible in practice. Thus (like for IS) one (or two) main directions are used for most applications.
- Disadvantage of OAS: Many strata (or an adaptive strata structure) are necessary for rare event simulations.
- Is it possible and sensble to combine IS and OAS to increase the variance reduction?

Stratified Importance Sampling: Single Estimate Case

- Stratified Importance Sampling (SIS): Applies OAS to an IS algorithm.
- Numerical results for IS of (Sak et al. 2010) and SIS for *t*-copula model with generalized hyperbolic marginals (parameters estimated from NYSE data)

	$P(R<t) \approx 0.05$						$P(R<t) \approx 0.001$					
	IS			SIS			IS			SIS		
d	VR	TM	ER	VR	TM	ER	VR	TM	ER	VR	TM	ER
2	6.1	0.23	6.5	296.5	0.30	243.5	183.5	0.21	201.7	4741.2	0.28	4046.9
5	8.4	0.79	8.4	110.3	0.90	96.8	278.5	0.79	247.3	3495.7	0.91	2684.9
10	5.3	1.59	5.0	11.4	1.75	9.8	66.6	1.64	60.5	198.3	1.91	154.1

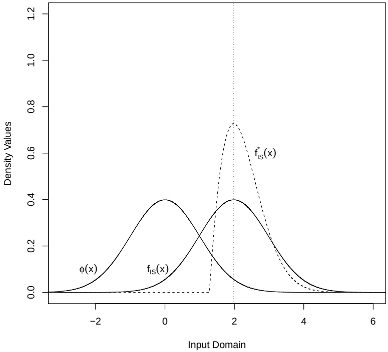
Variance reduction factors: $VR(\hat{x}) = V[\hat{x}_{NV}]/V[\hat{x}]$
Efficiency ratios: $ER(\hat{x}) = VR(\hat{x})\, TM[\hat{x}_{NV}]/TM[\hat{x}]$
TM exeuction time, $N \approx 100,000$ for all simulations.

Why works combination of IS and OAS so well?

- IS and OAS use the same direction and are very similar methods.
- We demonstrate their synergy effects for a one-dimensional example

simple example IS

$$y = E_{\phi}\left[q\left(Z\right)\right], \quad q\left(x\right) = \left\{e^{x} - 3.6\right\}^{+} \quad \text{and } Z \sim N(0,1).$$



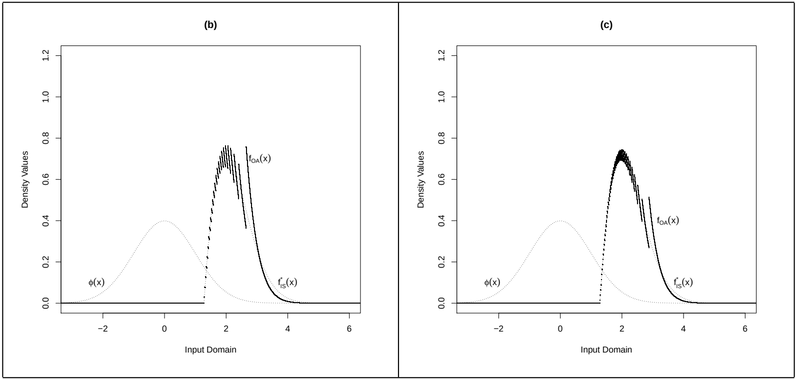
The original density $\phi\left(x\right)$,

the shifted IS density $f_{IS}\left(x\right)$ and

the optimal IS density $f_{IS}^{*}\left(x\right)$.

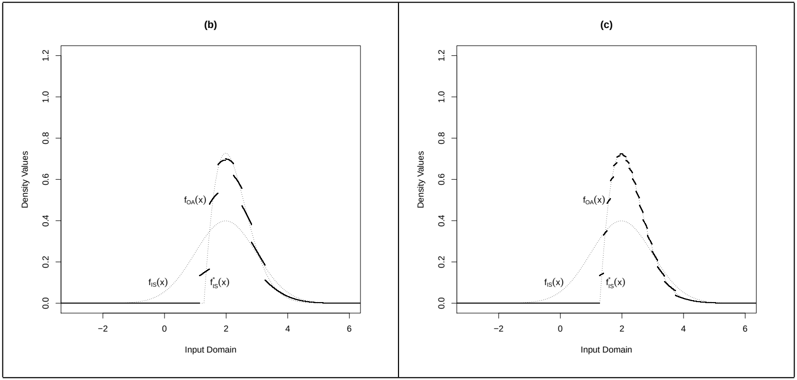
simple example with OAS

IS density corresponding to OAS with 100 and 500 strata.



simple example with SIS

IS density corresponding to SIS with 10 and 25 strata.



simple example: Variance Reduction Factors

Table: Comparison of Naive, IS, OAS 1000, and SIS 100.
exact solution 0.2815896024 .

	Estimate	Variance	VRF
Naive	0.27970	2.17E-05	1
IS	0.28192	3.77E-07	58
OAS 1000	0.28155	2.73E-09	7950
SIS 100	0.28159	8.53E-11	2.5e5

Advantages of combining IS and OAS

We have observed the following advantages when combining IS and OAS for risk simulations:

- IS helps that there are a smaller number of strata with return 0.
- Thus a smaller number of stratification intervals still leads to substantial variance reduction.
- IS helps stratification to obtain better estimates for the variances in the strata and thus better allocation fractions.

NEW IDEA

- Stratification can help to estimate many tailloss probabilities for different thresholds τ_j in a single simulation.
  That is important when estimating VaR and CVaR.

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Multiple Estimates Case

- Suppose we are interested in probability estimates for distinct threshold values, τ_j, $j = 1, \ldots, J$.
- We estimate the j-th tail loss probability and its relative error with the following formula:

$$\hat{x}_j = \sum_{i=1}^{I} p_i \hat{x}_{ij}, \qquad RE\left[\hat{x}_j\right] = \frac{\Phi^{-1}(0.975)}{\hat{x}_j \sqrt{N}} \sqrt{\sum_{i=1}^{I} \frac{p_i^2 \hat{\sigma}_{ij}^2}{\pi_i}}$$

- General questions
  - How should we define the "overall error" of our J estimation problems?
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Definition of Overall Error

$$\hat{x}_{ij} = \text{Prob}(Loss > \tau_j | X \in \xi_i)$$
$$\hat{x}_j = \text{Prob}(Loss > \tau_j) = \sum_{i=1}^{I} p_i \hat{x}_{ij}$$
$$\hat{s}_i^{jk} = Cov(\hat{x}_{ij}, \hat{x}_{ik}), \quad j, k = 1, \ldots, J$$

For the vector $\hat{\boldsymbol{x}}$ the variance-covariance matrix $\boldsymbol{\Sigma}$ depends on the allocation fractions:

$$\Sigma_{jk}(\boldsymbol{\pi}) = N^{-1} \sum_{i=1}^{I} \pi_i^{-1} p_i^2 \hat{s}_i^{jk} .$$

We define an overall error function:

$$\omega(\boldsymbol{\pi}) = g(\boldsymbol{\Sigma}(\boldsymbol{\pi})) .$$

Relevant Overall Error functions

- $\omega_{MSE}(\pi) = \sum_{j=1}^{J} \Sigma_{jj}(\pi)$, the mean squared error of all estimates,
- $\omega_{MSR}(\pi) = \sum_{j=1}^{J} \hat{x}_j^{-2} \Sigma_{jj}(\pi)$, the mean squared relative error of all estimates,
- $\omega_{MAXE}(\pi) = \max\{j : \Sigma_{jj}(\pi)\}$, the maximum of the squared errors of all estimates,
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mean squared relative error

Using the variances defined above we get:

$$\omega_{MSR}(\boldsymbol{\pi}) = \sum_{j=1}^{J} \hat{x}_j^{-2} \Sigma_{jj}(\boldsymbol{\pi}) = N^{-1} \sum_{i=1}^{I} \pi_i^{-1} p_i^2 \sum_{j=1}^{J} \hat{x}_j^{-2} \hat{s}_i^{jj},$$

We know from AOS that for a single estimate:

$$V\left[\hat{x}_j\right] = N^{-1} \sum\nolimits_{i=1}^{I} \pi_i^{-1} p_i^2 \sigma_i^2 \geq N^{-1} \left(\sum\nolimits_{i=1}^{I} p_i \sigma_i\right)^2.$$

It attains its lower bound for $\pi_i^* = \dfrac{p_i \sigma_i}{\sum_{k=1}^{I} p_k \sigma_k}$, $i = 1, \ldots, I$.

We can see that $\omega_{MSR}(\boldsymbol{\pi})$ has the same structure as $V\left[\hat{x}_j\right]$.

Replacing σ_i^2 by $\sum_{j=1}^{J} \hat{x}_j^{-2} \hat{s}_i^{jj}$ we thus can minimize $\omega_{MSR}(\boldsymbol{\pi})$.

mean squared relative error, cont.

$$\omega_{MSR}(\pi) \geq N^{-1}\left(\sum_{i=1}^{I} p_i \left(\sum_{j=1}^{J} \hat{x}_j^{-2}\hat{s}_i^{jj}\right)^{1/2}\right)^2$$

$\omega_{MSR}(\pi)$ attains its lower bound selecting

$$\pi_i^* = p_i\left(\sum_{j=1}^{J} \hat{x}_j^{-2}\hat{s}_i^{jj}\right)^{1/2} \Big/ \sum_{l=1}^{I} p_l \left(\sum_{j=1}^{J} \hat{x}_j^{-2}\hat{s}_i^{jj}\right)^{1/2}, \quad i = 1, \ldots, I.$$

This idea and the closed form solution can be generalized to all $\omega(\pi)$ that are linear functions of the s_i^{jk}.
(Theorem requires non-negativity condition. Assuming positive correlations is no problem for applications to simulation.)

Maximal relative error: Optimization Model

- New objective: Minimize the maximum relative error using the decision variables $\boldsymbol{\pi} = (\pi_1, \ldots, \pi_I)'$.
- We denote $\hat{a}_{ij} = \hat{x}_j^{-2} p_i^2 \hat{\sigma}_{ij}^2$ and add constraints which guarantee that the π_i are positive and sum to one.

$$\begin{aligned} \min \quad & \max_j \left\{ \sum_{i=1}^{I} \frac{\hat{a}_{ij}}{\pi_i} \right\} \\ \text{s.t.} \quad & \sum_{i=1}^{I} \pi_i = 1, \\ & \pi_i \geq 0, \qquad i = 1, \ldots, I. \end{aligned} \quad \Rightarrow \quad \begin{aligned} \min \quad & \omega \\ \text{s.t.} \quad & \omega - \sum_{i=1}^{I} \frac{\hat{a}_{ij}}{\pi_i} \geq 0, \quad j = 1, \ldots, J, \\ & \sum_{i=1}^{I} \pi_i = 1, \\ & \pi_i \geq 0, \qquad i = 1, \ldots, I. \end{aligned}$$

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- For the previous optimization model, we have developed a heuristic which searches for a suboptimal solution in the convex hull of the respective optimal solutions π^j, $j = 1, \ldots, J$.

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Numerical Results

- We consider $D = 5$ stocks under the t-copula model with Generalized Hyperbolic marginals and $J = 10$ equidistant threshold values.
- For IS we use a mixture of two densities.
- For SIS, we simply used a single IS density selected for the threshold $\tau^* = 0.75\tau_{max} + 0.25\tau_{min}$.
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Numerical Results cont.

- Following table shows the percentage relative errors and the variance reduction factors obtained under NV, IS and SIS. The execution times are 0.75, 0.95 and 1.25 seconds respectively.

		NV	IS		SIS	
τ	$\Pr(r(\mathbf{T}) \leq \tau)$	RE	RE	VR	RE	VR
0.895	0.0010	±19.89%	±2.65%	56.3	±0.77%	667.2
0.909	0.0014	±16.67%	±2.37%	49.5	±0.70%	567.1
0.917	0.0020	±14.13%	±2.12%	44.4	±0.69%	419.4
0.924	0.0028	±12.09%	±1.89%	40.9	±0.68%	316.1
0.932	0.0040	±9.89%	±1.77%	31.2	±0.66%	224.5
0.939	0.0060	±7.96%	±1.72%	21.4	±0.69%	133.1
0.947	0.0094	±6.46%	±1.82%	12.6	±0.66%	95.8
0.954	0.0154	±4.91%	±2.14%	5.3	±0.65%	57.1
0.962	0.0267	±3.75%	±2.26%	2.8	±0.64%	34.3
0.973	0.0500	±2.69%	±9.23%	0.1	±0.69%	15.2

Estimating Conditional Excess

- Simulating the conditional excess: $E\left[Loss|Loss \geq \tau\right]$ requires a ratio estimate; this makes variance reduction more difficult.
- Literature: "Use the same IS density as for tail-loss probabilities."
- The variance of the ratio estimate: $V\left[\hat{x}_1/\hat{x}_2\right] \approx \hat{x}_1^2 \hat{x}_2^{-4} \Sigma_{22}\left(\boldsymbol{\pi}\right) - 2\hat{x}_1 \hat{x}_2^{-3} \Sigma_{12}\left(\boldsymbol{\pi}\right) + \hat{x}_2^{-2} \Sigma_{11}\left(\boldsymbol{\pi}\right) = \omega\left(\boldsymbol{\pi}\right).$
- To reach optimal allocation for stratification we can use the theorem above to minimize the variance of a ratio estimate.

For Conditional Excess we obtained stratification and stratified IS algorithms with optimal allocation for:

- a single threshold
- for several thresholds minimizing the mean squared relative error.

Conclusions

- For practically relevant examples SIS (combination of IS and stratification) increases the efficiency of tail loss probability estimates under the t-copula model.
- Compared to the methods in the literature, the variance of the estimates are substantially reduced without a significant increase in the execution time.
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- The mean square (relative) error of all estimates can be minimized using the variance estimates of a pilot run and a simple closed form formula for the allocation fractions.
- To minimize the maximal squared (relative) error of all estimates we have developed a fast and simple heuristic to find close to optimal allocation fractions.
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Questions?

THANK YOU !

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